

Social inequality in mortality among adults and elderly in northern Sweden 1851–2013¹

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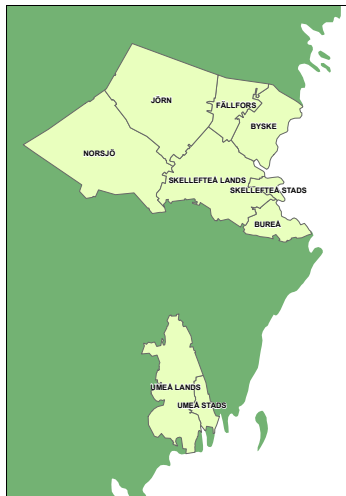
Questions

1. Has **inequality in mortality** between **social classes** increased in the adult and elderly population?
2. Are there any **gender differences** in the effect of social position?
3. Is social position **equally important** among the **retired** population as for those in **working age**?
4. Is there a difference between classes in mortality from **cardiovascular** diseases and **cancers** respectively?
5. Are there large differences in survival depending on level of **education** and **income** categories?

Outline

1. Data management
2. Modeling considerations
3. Analysis
4. Results
5. Alternative Results

Here!



Skellefteå:

- Small town, heavy industry in the nineteenth century.

Umeå:

- Not so small town, older. Center for administration, education, and military.

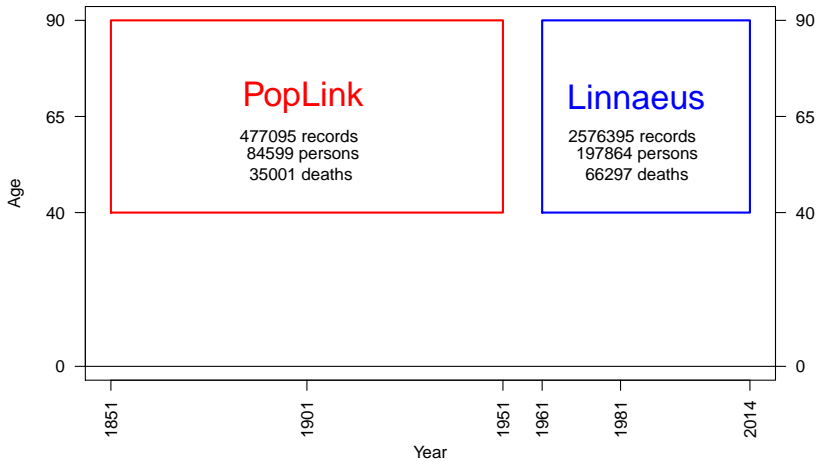
Rural areas:

- Very similar: Small farms dominating.

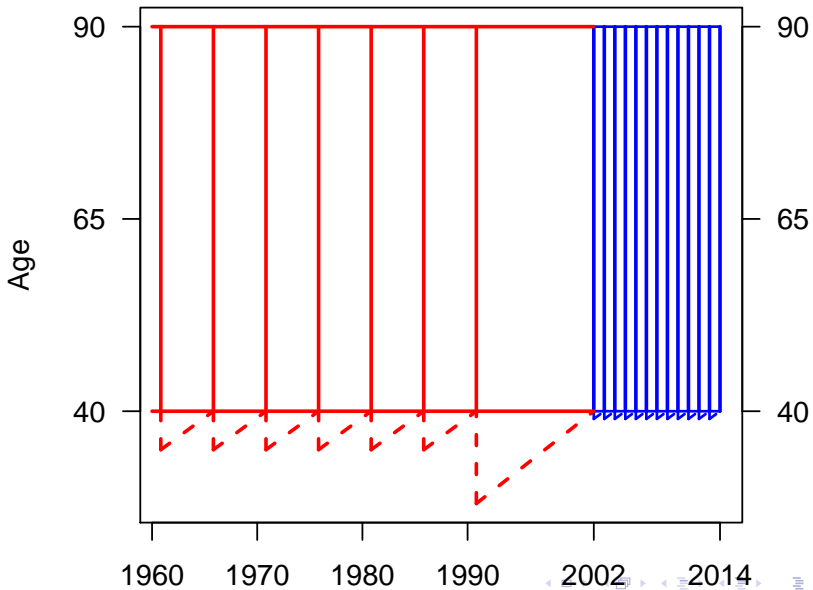
Data sources

- **Swedish church book** records (1851–1950).
 - digitized by the **Demographic Data Base** at Umeå University,
 - **POPLINK**
 - <http://www.cedar.umu.se/>
- The **Linnaeus** data base (1960–2013)
 - Census data 1960, 1965, . . . , 1990
 - Population registers (**LISA**, 1986–2013)
 - **Income**
 - **Education**
 - Death information from **National Board of Health and Welfare**
 - Death dates from 1 January 1961 to 31 December 2013.
 - **Causes of death** (**ULORSAK**, “main cause of death”)

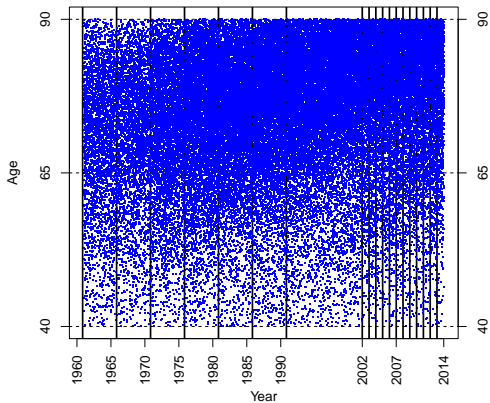
The sampling frame



The FOB and Lisa data



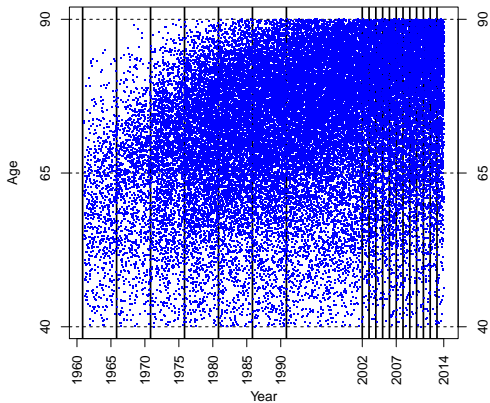
The Linnaeus data with deaths



	Number	Deaths
Women	97988	30275
Men	99876	36022
Total	197864	66297

Note: Persons with missing HISCLASS (occupation) are included.

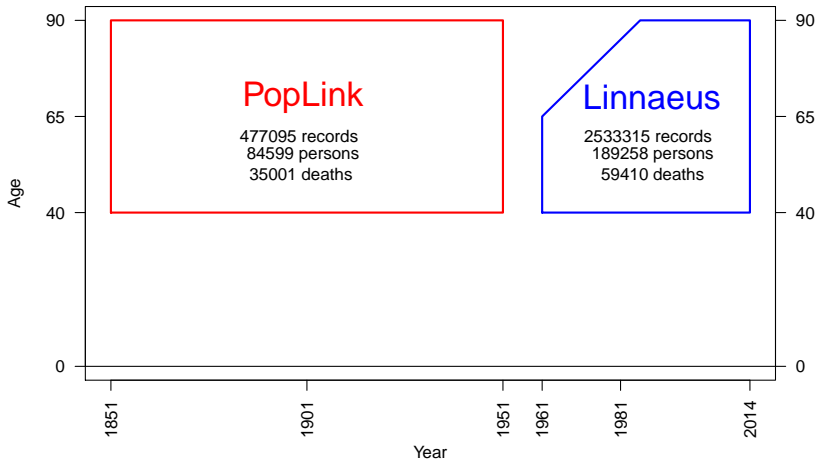
The Linnaeus data with deaths



	Number	Deaths
Women	70249	13888
Men	86427	28991
Total	156676	42879

Note: Persons with **missing HISCLASS** (occupation) are excluded.

The sampling frame, II



Variables

sex

birth date By quarter (15 Feb, May, Aug, Nov)

death date By day ([Socialstyrelsen](#))

cause of death from [Socialstyrelsen](#)

parish Urban/Rural

HISCLASS elite, middle, worker.

income 1990–2005, grouped by quartiles.

education 1990–2005, seven categories.

period 1851–1875, . . . , 2008–2013 (10 periods).

cohort 1761–1771, . . . , 1971–1974 (22 birth cohorts).

HISCLASS

1. **elite**, HISCLASS 1 and 2. Higher managers and professionals.
2. **middle**, HISCLASS 3, 4, 5, 6, and 8. Lower managers, lower white collar, and farmers.
3. **worker**, HISCLASS 7, 9, 10, 11 and 12. Workers of different skills including farm workers.

Preliminary Cox regression

```
library(eha)
system.time(fit <- coxph(Surv(enter, exit, event) ~
                        hisclass * period + civst + urban + sex,
                        data = vb0))

   user  system elapsed
151.360   0.444  151.806

system.time(dr <- drop1(fit, test = "Chisq"))

   user  system elapsed
455.236   0.944  456.188
```

It takes time (more than 10 minutes)!

ANOVA

	Df	AIC	LRT	Pr(>Chi)
<none>		2227937.83		
civst	2	2228508.16	574.34	0.0000
urban	1	2227962.96	27.13	0.0000
sex	1	2231526.62	3590.79	0.0000
hisclass:period	27	2228575.58	691.75	0.0000

And with so much data, everything tends to be **statistically** significant.

Test of Proportional Hazards

user	system	elapsed
1063.928	0.916	1064.028

Another 19 minutes!

Proportional hazards, results

```
#system.time(cz <- cox.zph(fit))
```

```
cz$table[NROW(cz$table), ]
```

rho	chisq	p
NA	2064.07	0.0000

- As usual, **highly significant** result (PH assumption violated).
- Should we worry?
 - Yes
 - No

Sufficient statistics

Aggregate:

```
vb0Tab <- aggregate(vb0[, c("event", "exposure")],  
                    vb0[, c("sex", "age", "period", "cohort",  
                            "urban", "civst", "hisclass")],  
                    FUN = sum)  
vb0Tab$age <- factor(vb0Tab$age)  
nr <- NROW(vb0Tab)
```

- Gives a table assuming **piecewise constant hazards** on five-year intervals: 40–45, 45–50, . . . , 85–90 ("age").
- Number of cells: $2 \times 10 \times 10 \times 20 \times 2 \times 3 \times 4 = 96000$
- Number of **non-empty** cells: **10785**.

<http://capa.ddb.umu.se/ds/aggregate.html>

PH regression with the table

- Piecewise constant hazard model generates the same likelihood as a Poisson model.
 - Does **not** imply that the number of deaths follow a Poisson distribution!

```
system.time(fitp <- glm(event ~ offset(log(exposure)) + age +
                        hisclass * period + civst + urban + sex,
                        data = vb0Tab, family = poisson))
```

```
user  system elapsed
0.884  2.216   0.570
```

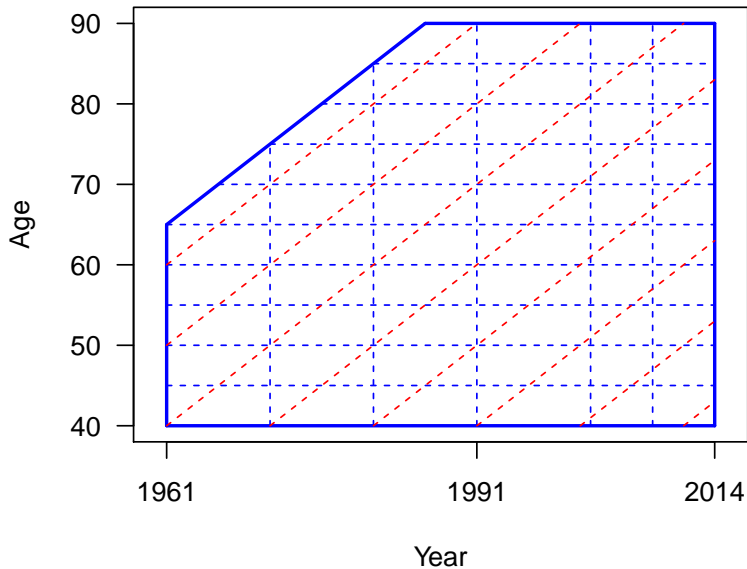
```
system.time(dr <- drop1(fitp, test = "Chisq"))
```

```
user  system elapsed
2.476  6.440   1.164
```

ANOVA with the table

	Df	Deviance	AIC	LRT	Pr(>Chi)
<none>		17539.79	43607.22		
age	9	116442.66	142492.08	98902.87	0.0000
civst	2	18254.46	44317.88	714.67	0.0000
urban	1	17553.35	43618.78	13.56	0.0002
sex	1	21008.66	47074.08	3468.87	0.0000
hisclass:period	27	18167.89	44181.31	628.09	0.0000

Age-Period-Cohort data



Counting "cells"

- There are 115 cells in one figure
- There are $2 \times 2 \times 3 \times 4 = 48$ such figures.
- In total: $115 \times 48 = 5520$ combinations.
- Number of non-empty combinations: 4072.

	sex	age	period	cohort	urban	civst	hisclass	event	exposure
158	female	70	1981-1990	1901	FALSE	unmarried	elite	0	5.3750000
159	male	75	1981-1990	1901	FALSE	unmarried	elite	0	6.5000000
160	female	75	1981-1990	1901	FALSE	unmarried	elite	2	11.6873370
161	male	80	1981-1990	1901	FALSE	unmarried	elite	2	1.8648946
162	female	80	1981-1990	1901	FALSE	unmarried	elite	0	6.3761636
163	female	85	1981-1990	1901	FALSE	unmarried	elite	0	0.3755818

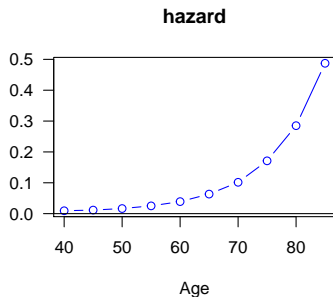
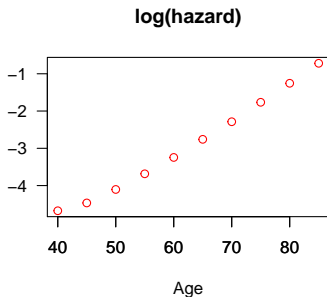
An **anonymous** data set!

The 'baseline hazard function'

	Estimate	Std. Error
(Intercept)	-4.68	0.13
age45	0.21	0.02
age50	0.57	0.02
age55	0.99	0.02
age60	1.43	0.02
age65	1.92	0.02
age70	2.39	0.02
age75	2.91	0.02
age80	3.42	0.02
age85	3.96	0.02

- (Intercept) is log hazard in 40–45. Reference interval.
- age45 is the difference in log hazard between 45–50 and 40–45,
- age50 is the difference in log hazard between 50–55 and 40–45,
- Etcetera ...

Baseline hazards



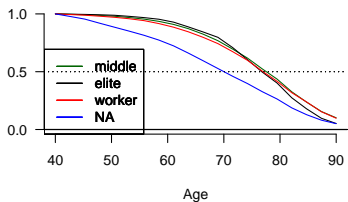
Why not try:

$$\log\{h(t)\} = \alpha + \beta t, \quad 40 < t \leq 90,$$

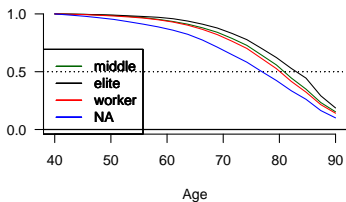
a **Gomperz** distribution?

Survival analysis, men 1961-2013

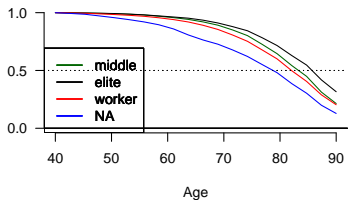
Men 1981-1990



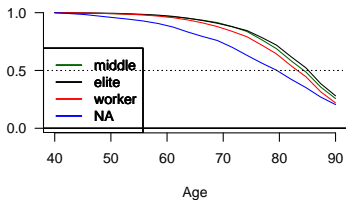
Men 1991-2001



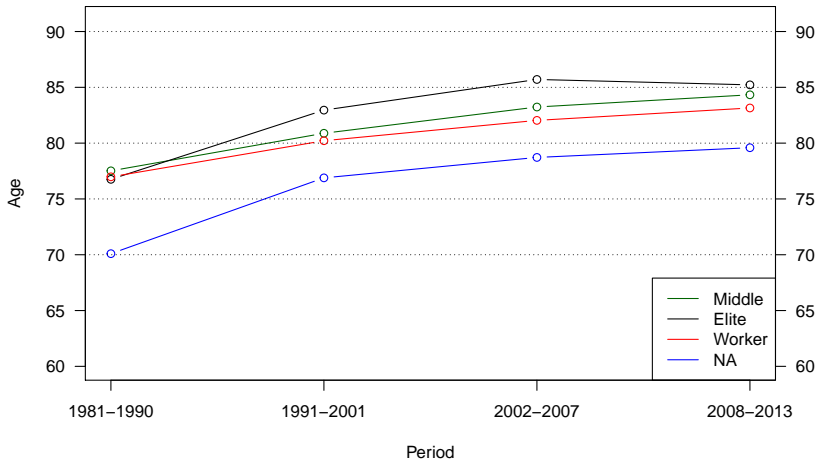
Men 2002-2007



Men 2008-2013

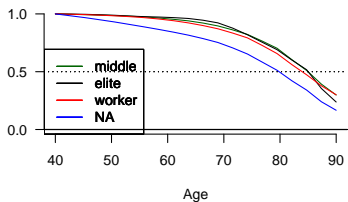


Median life after 40, men 1961-2013

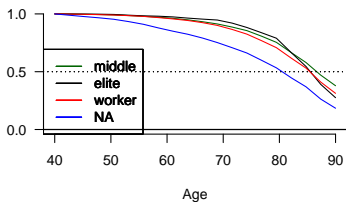


Survival analysis, women 1961-2013

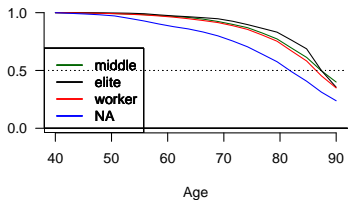
Women 1981-1990



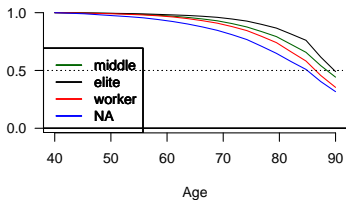
Women 1991-2001



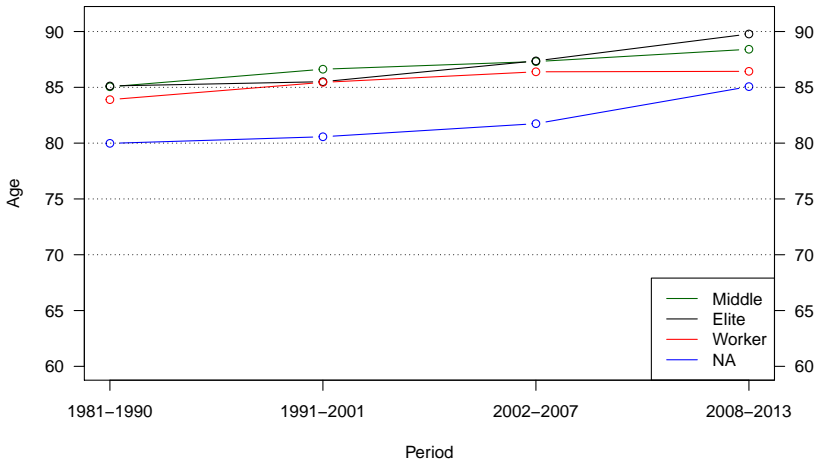
Women 2002-2007



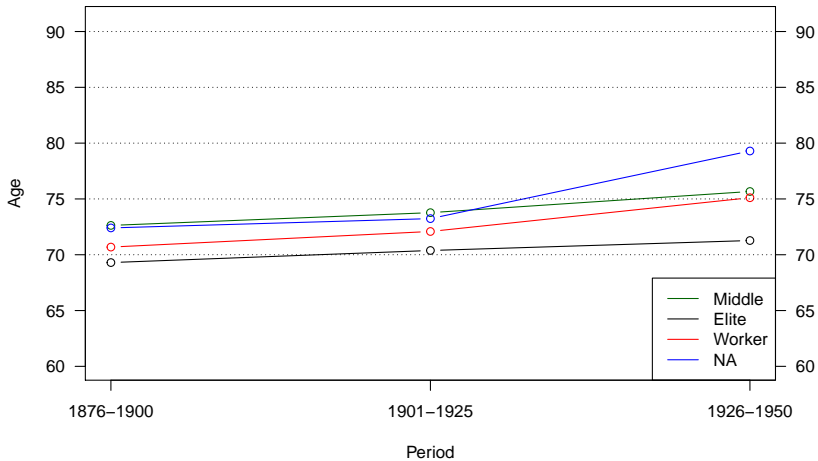
Women 2008-2013



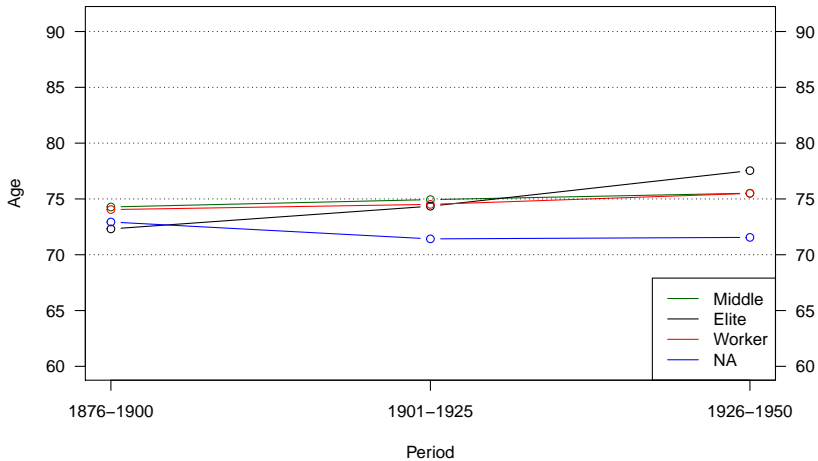
Median life after 40, women 1961-2013



Median life after 40, men 1851-1950

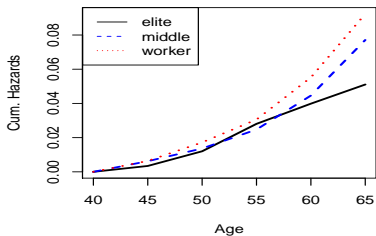


Median life after 40, women 1851-1950

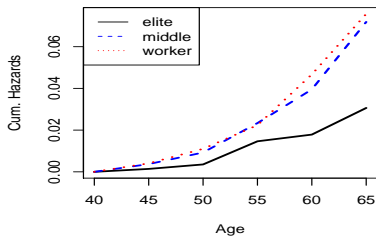


Proportional hazards for women?

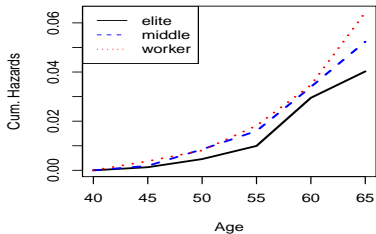
1981–1990



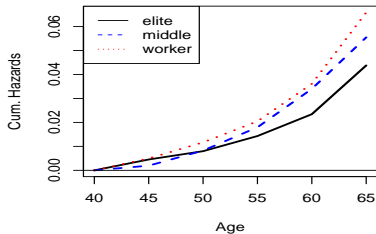
1991–2001



2002–2007

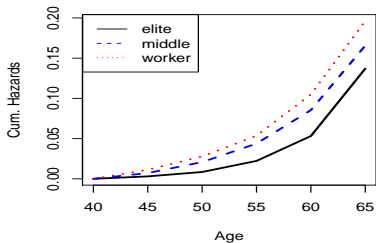


2008–2013

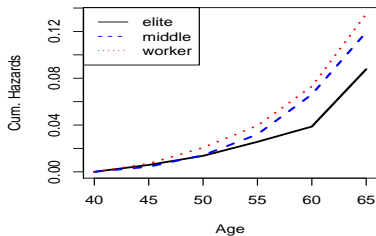


Proportional hazards for men?

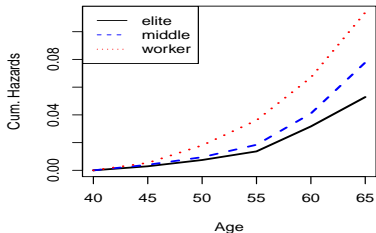
1981–1990



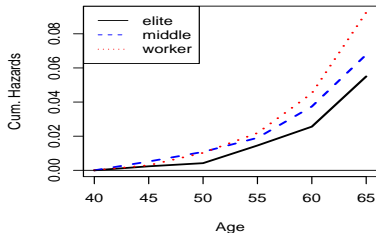
1991–2001



2002–2007



2008–2013



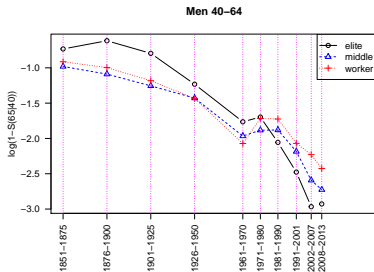
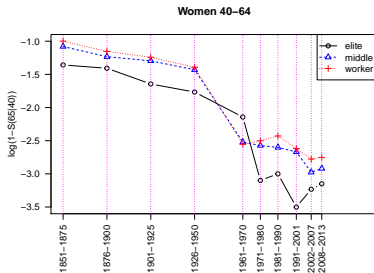
No proportional hazards (wrt HISCLASS)!

- Fit each HISCLASS *separately*, for each time period.
- Use the *cumulative hazard* at 65 (or 90) for comparison.

So, we fit separate PH models for each combination of

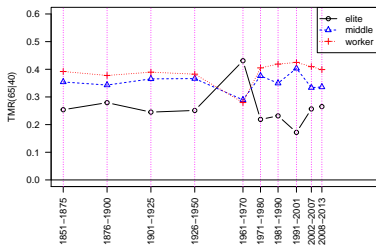
- *sex*
- *period*: 1851–75, 1876–1900, 1901–10, etc.
- *hisclass*
- *age group*: 40–64, 65–89.
- *cause of death*

Women and men, age 40–64

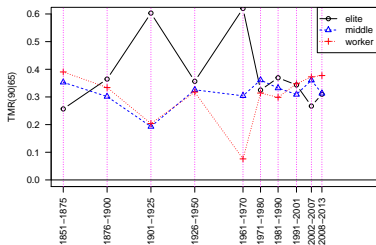


Relative differences, ages 40–64 and 65–89

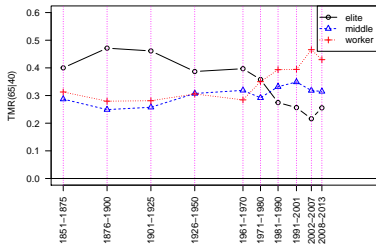
Women 40–64



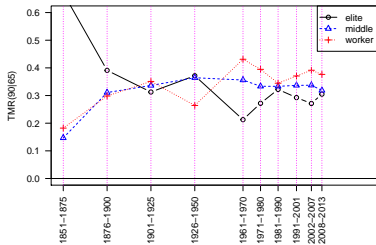
Women 65–89



Men 40–64

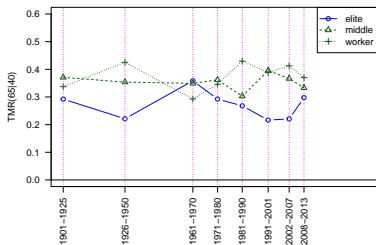


Men 65–89

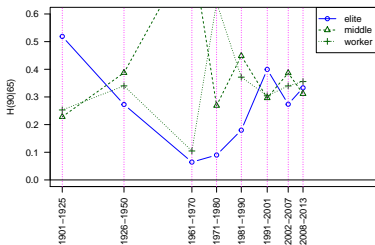


Cancer mortality, ages 40-64 and 65-89

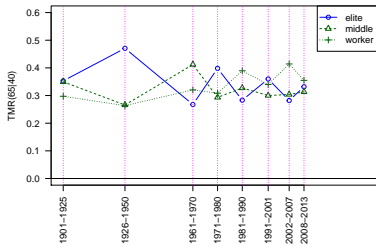
Women 40-64



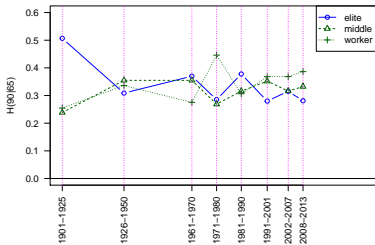
Women 65-89



Men 40-64

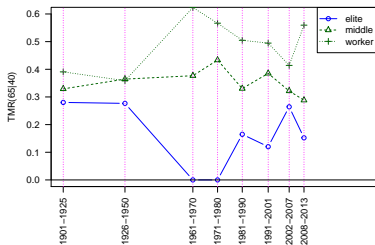


Men 65-89

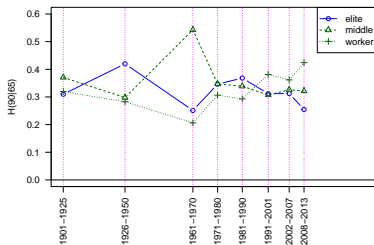


Cardiovascular mortality

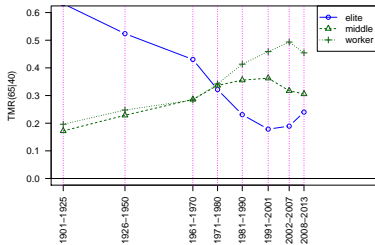
Women 40-64



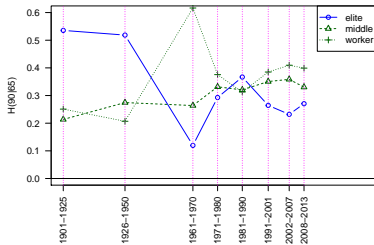
Women 65-89



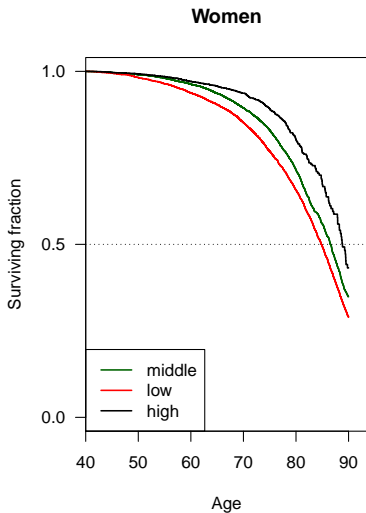
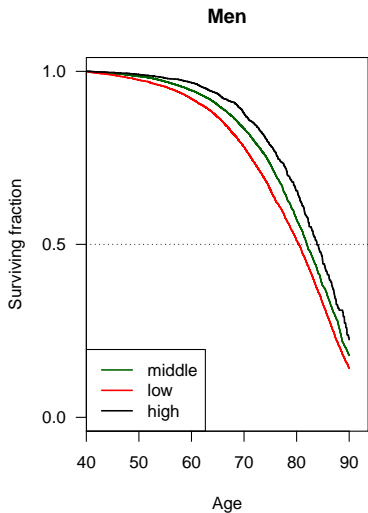
Men 40-64



Men 65-89

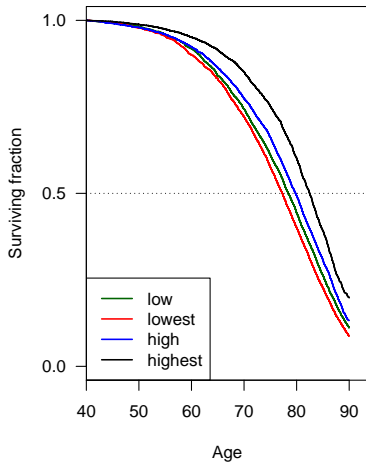


Education 1990–2013

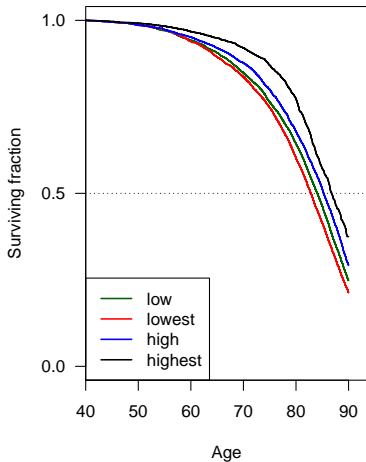


Income 1986–2005

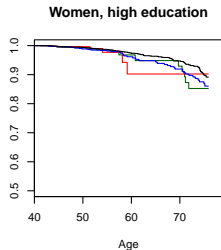
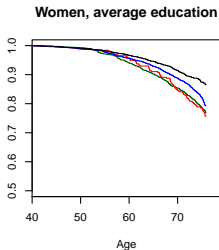
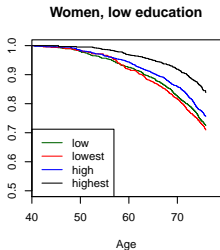
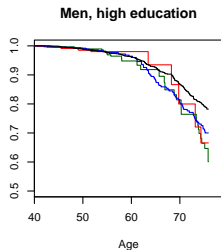
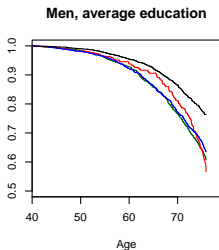
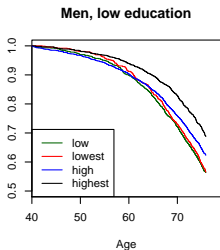
Men



Women



Income by education, 1990–2005, ages 40–75!



Answers(?)

1. Has **inequality in mortality** between **social classes** increased?
 - Not really
2. Are there any **gender differences** in the effect of social position?
 - Yes
3. Is social position **equally important** among the **retired** population as for those in **working age**?
 - Not really
4. Is there a difference between classes in mortality from **cardiovascular** diseases and **cancers** respectively?
 - Yes
5. Are there large differences in survival depending on level of **education** and **income** categories?
 - Yes, definitely.

Conclusion

- Improved survival characterised all HISCLASSES.
- There are gender differences in the social pattern
- No difference in the social pattern in the two studied age groups.
- The upper class(es) gradually take over as the most healthy group.
- Education and income are the most important variables during late time period.